

$$\sigma_{\epsilon}^2 = \frac{1}{N-d} \sum_{k=1}^N (y^{(k)} - \phi^{(k)\top \theta})^2$$

$$\text{cov}(\hat{\theta}) = \sigma_{\epsilon}^2 (\phi_N^{\top} \phi_N)^{-1}$$

CH 5: MAXIMUM LIKELIHOOD AND BAYESIAN ESTIMATION

MODEL:

$$\begin{aligned} \eta: \mathbb{R}^d &\rightarrow \mathbb{R}^N \\ y, \varepsilon &\in \mathbb{R}^N \end{aligned}$$

$$Y = M(\theta) + \varepsilon$$

$$Y = \begin{pmatrix} y_1 \\ \vdots \\ y_N \end{pmatrix} \quad \varepsilon = \begin{pmatrix} \varepsilon_1 \\ \vdots \\ \varepsilon_N \end{pmatrix}$$

$M(\theta)$ DETERMINISTIC, ε ZERO MEAN, ...
(INDEPENDENT)
(I.I.D.)

5.1 MAX. LIKELIHOOD (ML)

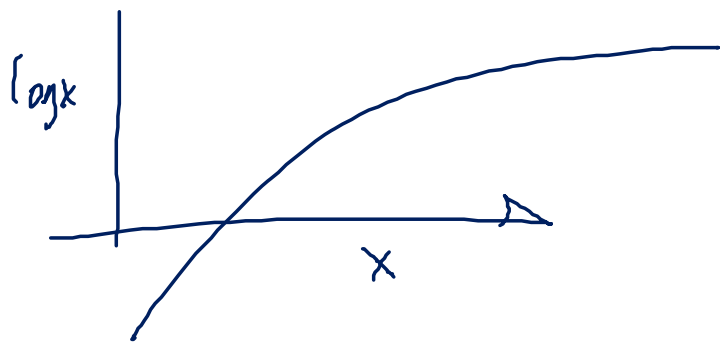
DEF: $L(\theta) = p(y|\theta)$ LIKELIHOOD FUNCTION
"PDF OF Y GIVEN θ FOR Y FIXED"

DEF: ML-ESTIMATOR: GIVEN Y , MAXIMIZE $p(Y|\theta)$ WRT. θ

EX: $\epsilon_i, \epsilon_1, \dots, \epsilon_N$ ~~IID~~. ZERO MEAN, INDEPENDENT GAUSSIAN (σ_i^2)

$$p(Y|\theta) = \prod_{i=1}^N p(y_i|\theta) = C \prod_{i=1}^N \exp\left(-\frac{(y_i - \mu_i(\theta))^2}{2\sigma_i^2}\right)$$

$$-\log p(Y|\theta) = -\log C + \sum_{i=1}^N \frac{(y_i - \mu_i(\theta))^2}{2\sigma_i^2}$$



$$\arg \max p(Y|\theta) = \arg \min -\log p(Y|\theta)$$

MAX. LIKELIHOOD $\hat{=}$ NONLINEAR LEAST SQUARES

$$\sum_{i=1}^N \frac{(y_i - \mu_i(\theta))^2}{2\sigma_i^2} = \frac{1}{2} \| S^{-1} (y - M(\theta)) \|_2^2$$

$$S = \begin{pmatrix} \sigma_1 & & \\ & \ddots & \\ & & \sigma_N \end{pmatrix} = \frac{1}{2} \| \underbrace{S^{-1} y - S^{-1} \cdot M(\theta)} \|_2^2$$

$$S = \sum_{i=1}^N \frac{1}{2\sigma_i^2} = \frac{1}{2} \| F(\theta) \|_2^2$$

→ SOLVE WITH "lsqnonlin"
IN MATLAB (FINDS LOCAL
MINIMUM)
(GAUSS-NEWTON METHOD)

MORE GENERAL, CAN TREAT CORRELATED
NOISE $\epsilon \sim \mathcal{N}(0, \Sigma)$

$$p(y|\theta) = C \cdot \exp\left(-\frac{(y - \mu(\theta))^T \Sigma^{-1} (y - \mu(\theta))}{2}\right)$$

$$-\log p(y|\theta) = \text{const} + \underbrace{\frac{1}{2} (y - \mu(\theta))^T \Sigma^{-1} (y - \mu(\theta))}_{= \frac{1}{2} \| \Sigma^{-\frac{1}{2}} y - \Sigma^{-\frac{1}{2}} \mu(\theta) \|_2^2}$$

REGULARIZATION: ADDITION OF AD-HEC / A-PRIORI KNOWLEDGE CAN HELP TO REDUCE NON-UNIQUENESS

$$\min \frac{1}{2} \|y - M(\theta)\|_2^2 + \frac{\alpha}{2} \|\theta - \bar{\theta}\|_2^2 \quad \text{WITH } \alpha \text{ SMALL}$$

"PSEUDO-MEASUREMENT"
 (→ BAYESIAN ESTIMATION)

5.1.1 L1-ESTIMATION

$$\sum_{i=1}^N \frac{|y_i - \mu_i(\theta)|}{2 a_i}$$

"ROBUST ESTIMATION"

NON-SMOOTH

(LAPLACE)

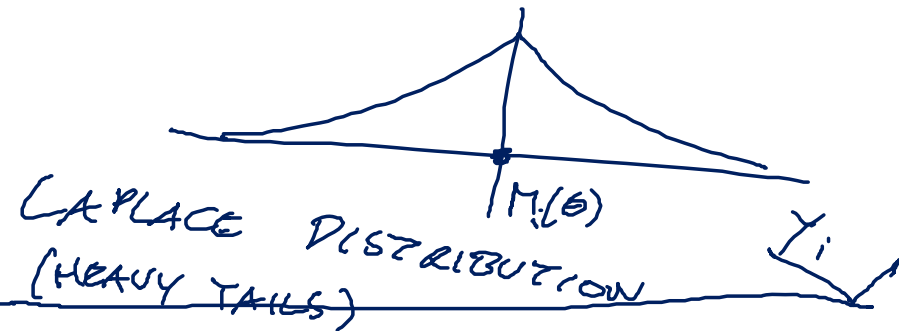
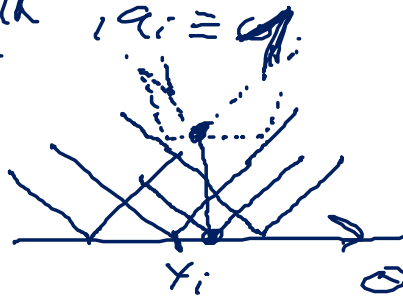
$$p(y_i | \theta) = C \cdot \exp\left(-\frac{|y_i - \mu_i(\theta)|}{2 a_i}\right)$$

EXAMPLE: L1

$$\mu_i(\theta) = \theta \in \mathbb{R} \quad a_i = \sigma$$

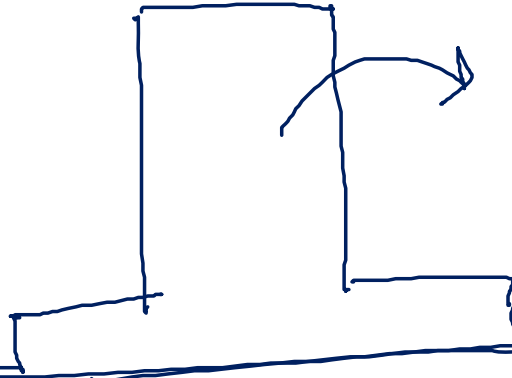
$$\arg \min_{\theta} \sum |y_i - \theta| = \text{MEDIAN}$$

(ROBUST AGAINST OUTLIERS)



READING SCRIPT

NOT STARTED YET	CH. 1 (INTERO READ, BUT NOT CH2)	CH 2 (PRODS)	CH 3 (OPT)	CH 4 (LLS)	CH 5	MORE
1	1	4	21	4		0



WORKLOAD PER WEEK (IN ADDITION TO PRESENCE IN LECT. & EX.)
 (AVERAGE UNTIL NOW)

1h	2h	4h	6h	8h	10 or more
0	0	2	6	13	15



WHAT IS TIME USED FOR (50% OR MORE)

EXERCISE SHEETS	READING SCRIPT
MAJORITY	~ 1/3

MATLAB	VARIED
13	~20

OTHER SOURCES

PROG. & STATISTICS TEXTBOOK

5.2 BAYESIAN ESTIMATION

INSTEAD OF $p(y|\theta)$ WE MAY WANT \rightarrow

MAXIMIZE
" θ GIVEN y "

$p(\theta|y)$

"BAYES FORMULA"

ML-TERM
A-PRIORI

$$p(\theta|y) = \frac{p(\theta, y)}{p(y)} = \frac{p(y|\theta) \cdot p(\theta)}{p(y)}$$

CONST.
WRT.
 θ

DEF:

"MAXIMIZE $p(\theta|y)$ " $\hat{=}$ "MAXIMUM A-POSTERIORI (MAP) ESTIMATION"

$$\arg \max_{\theta} p(\theta|y) = \arg \min_{\theta} (-\log p(\theta|y)) = \arg \min_{\theta} \left(\underbrace{-\log p(y|\theta)}_{\text{ML-ESTIMATION}} - \underbrace{\log p(\theta)}_{\text{"REGULARIZATION" A-PRIORI INFO.}} \right)$$

NOTE: A-PRIORI INFORMATION LEADS \rightarrow BIAS

RELATION BETWEEN ML & MAP

ML IS SPECIAL CASE OF MAP, BUT WITH ZERO A-PRIORI INFORMATION

MAP CAN BE CONS. SPECIAL CASE OF ML, BUT WITH "PSEUDO-MEASUREMENT"

$-\log p(y \theta)$	$-\log p(\theta)$
ML	
	MAP